

# MALNUTRITION POINT-PREVALENCE FROM 2012 TO 2019 AND ASSOCIATED HEALTH-OUTCOMES IN ADULT PATIENTS IN RURAL HOSPITALS

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**Abstract:** *Background:* Malnutrition negatively impacts hospitalised patients and the healthcare system. *Objectives:* 1) report point-prevalence of hospital malnutrition from 2012 to 2019; and 2) determine if there was an association between nutrition status and health-related outcomes. *Design:* Point-prevalence of malnutrition was determined by three (2012, 2014, and 2019) cross-sectional studies. Health-related outcomes, assessed by a prospective cohort study in 2014, were length of stay, in-hospital mortality, hospital readmission, infection, falls, fractures, and pressure wounds. *Setting:* three Australian rural hospitals. *Participants:* Adult inpatients. *Measurements:* Nutrition status was assessed with the Subjective Global Assessment (SGA) tool. *Results:* Malnutrition point prevalence was 39% in 2012 (n=62), 48% in 2014 (n=128), and 28% in 2019 (n=96); where the prevalence in 2019 was significantly lower than in 2014 ( $p<0.017$ ). The 2019 (median age 70 years) sample was younger than the 2012 (median age 80 years) and 2014 (median age 78 years) samples ( $p<0.05$ ). Mortality and falls rate were higher in the severely malnourished participants ( $p<0.05$ ); and severe malnutrition may predict mortality (Adjusted OR: 3.47 (95%CI: 0.94, 12.78]  $p=0.061$ ). *Conclusions:* Nutrition status did not predict other health-related outcomes. The rate of malnutrition in rural hospitals was consistently high and may increase the risk of in-hospital mortality.

**Key words:** Malnutrition, hospitals, nutrition assessment, subjective global assessment, mortality.

## Introduction

Protein-energy malnutrition (herein referred to as 'malnutrition') negatively impacts the patient and healthcare system alike (2, 3), a major concern as the prevalence has been reported internationally at 30-50% across inpatient and residential settings, and 1-25% across community settings (4-8). Malnutrition is the unintended loss of lean mass (muscle, immune and blood cells, viscera), with or without fat loss, due to inadequate intake, uptake, and/or utilisation of protein and energy to meet requirements (4, 9). Older adults are at greater risk of malnutrition due to their susceptibility of aetiological factors including psychological, socio-economic, and physiological changes and an overall increase in multi-

morbidities and polypharmacy (9). A consequence of malnutrition is further morbidity, requiring increased healthcare resources including but not limited to hospital beds, multidisciplinary staff, and pharmaceutical and nutritional medicine (10). In particular, malnutrition increases risk of infection, pressure ulcers, poor wound healing, decreased response to medical treatment and pharmaceuticals, decreased respiratory function, decreased muscle repair, and overall functional impairment; leading to decreased quality of life and increased risk of mortality (3, 6, 11).

While several large-scale studies have reported the prevalence and outcomes of hospital malnutrition, the rural context requires specific examination as populations in rural areas are ageing more rapidly than in urban areas (11, 15-17). Rural areas face increased challenges in providing health and aged care due to the higher cost of establishing and delivering services, the limited availability of and access to health professionals, and less availability of informal care networks (12-14). Not only is access to health care more limited in rural areas, rural-dwelling older adults are also more in need of health and aged care services. A recent meta-analysis and meta-regression of international data found the prevalence of malnutrition in rural-dwelling older adults living at home was double that of urban-dwelling older

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adults (5). Therefore, the prevalence and health-related outcomes of malnutrition in rural hospitals is of interest, so that policies may appropriately support patients in the continuum of care from hospital to home or residential care.

## Research aims

In adult patients admitted to three rural hospitals in Australia, the aims of this study were to: 1) report point-prevalence of malnutrition from 2012 to 2019; and 2) determine if there was an association between nutrition status and health-related outcomes.

## Materials and Methods

### Study design

The point-prevalence of malnutrition was assessed using three cross-sectional studies conducted in 2012, 2014, and 2019. The association between malnutrition and health-related outcomes was evaluated using a prospective observational study in 2014. Participants gave their verbal consent to participate in the study. The project was approved by the Human Research Ethics Committees in April 2018 (QA249) as a quality assurance project. This study has been reported according to the STROBE Statement for cohort studies (15) and was retrospectively registered with ANZCTR (ACTRN12619000342112) (19).

### Setting and sample

All three hospitals within a rural government-funded local health district in northern New South Wales, Australia, were conveniently sampled in 2012. Reflecting the staffing resources available for each cross-sectional study, the medical, surgical, general (not diagnosis or treatment specific), and/or rehabilitation wards were sampled (Table 1).

**Table 1**

Rural hospitals and wards sampled by the three cross-sectional studies

Sample	Hospital	Wards included
2012	Hospital 1	Medical, surgical, rehabilitation
	Hospital 2	
	Hospital 3	
2014*	Hospital 1	Medical, general, surgical, rehabilitation
	Hospital 2	
	Hospital 3	
2019	Hospital 1	Medical, general, surgical

\* This sample was also that used in the prospective observational study.

The prospective cohort study was conducted using the 2014 sample due to availability of data and its larger sample size. Patients were eligible if they were 18 years or older and were admitted as inpatients to study sites during the recruitment phase of one to two weeks. No exclusion criteria were applied.

### Participant characteristics and potentially confounding variables

Participant characteristics of age (years) and sex (male/female) were recorded for all participants. The 2014 sample were also described by comorbidities and medications. The number of active comorbidities were categorised into medical diagnostic groups: cancer, digestive, musculoskeletal, circulatory, respiratory, nervous, skin, reproductive, kidney, infectious, endocrine, injuries, ear, blood, and other. The number of medications were categorised into 24 drug classes based on the profile of medications recorded from the cohort (16).

### Outcomes

Malnutrition was determined by the SGA tool which rates patients as A = well nourished, B = mild-moderate malnutrition, or C = severe malnutrition (17, 18). The presence of malnutrition was the primary outcome to answer the first research question (point-prevalence) and the independent variable to answer the second research question (health-related outcomes).

The primary health-related outcome was length of hospital stay, defined by the number of days including the day of admission and discharge. Secondary health-related outcomes were in-hospital mortality (yes/no), hospital readmission (yes/no), in-hospital fall (yes/no), fall in subsequent hospital admissions (yes/no), pressure ulcer (yes/no), fracture acquired in hospital (yes/no), urinary tract or respiratory tract infection (yes/no). Health-related outcomes were measured from the day of hospital admission to three months post-discharge.

### Data Collection

In 2012 and 2014, nutrition status was assessed by department dietitians over a 7-day period within each ward using the SGA. In 2019, nutrition status was assessed by one student-dietitian (EL) using the SGA over two weeks. Only the student dietitian in 2019 received training in correct SGA use, whereas department dietitians were expected to be competent in nutrition assessment due to years of experience. The SGA is comprised of two main components: medical and physical assessment. Changes in weight, dietary intake, gastrointestinal symptoms, and nutrition-related functional capacity were observed from a combination of patient records and patient interview for the medical

**Table 2**  
Age, sex, and malnutrition point-prevalence of the 2012, 2014, and 2019 participant samples

Cohort year	2012 n = 62	2014 n = 128	2019 n = 96	p-value
Age years (median (IQR))	80 (60-85)	78 (65-86)	70 (58-82)	<0.001, 0.015*
Sex (n (%))				0.903†
- Males	28 (45)	60 (47)	51 (53)	
- Females	34 (54)	68 (53)	45 (47)	
Malnourished (n (%))				
-Total‡	24 (39)	62 (48)	27 (28)	<0.001§
-Males	12 (19)	31, (24)	13 (14)	0.794
- Females	12 (19)	31 (24)	14 (15)	

\* Kruskal-Wallis H Test applied followed by Mann-Whitney U tests to determine significance between cohort years. The Bonferroni correction was applied at 0.05 level. Cut-off value for significance for Mann-Whitney U tests at  $p=0.017$  (0.05/3) for 3 pairwise comparisons between cohorts. Significance for age found between pairs 2012 vs 2019 and 2014 vs 2019 cohorts; † No significance found between sex across cohorts; ‡ Malnourished = B (mild-moderate malnutrition) and C (severe malnutrition) rating combined; § Kruskal-Wallis H Test applied followed by Mann-Whitney U tests to determine significance between cohort years. The Bonferroni correction was applied at 0.05 level. Cut-off value for significance for Mann-Whitney U tests at  $p=0.017$  (0.05/3) for 3 pairwise comparisons between cohorts. Significance for malnourished patients only found between 2014- and 2019-year groups; || No significance found for sex effect on malnutrition across cohorts.

component; while evidence of oedema, ascites, and loss of subcutaneous fat and muscle was assessed during a physical examination to inform the physical component. A patient who demonstrated negative changes to their oral intake and failed to meet their nutritional requirements with evidence of muscle and fat deterioration were classified as exhibiting a degree of malnutrition. Participant characteristics and the health-related outcomes were observed from the medical record.

### Data Analysis

Data analysis was completed using IBM SPSS Statistics 25 [IBM Corp. Released 2017. IBM SPSS Statistics for Windows, Version 25.0. Armonk, NY: IBM Corp]. Descriptive statistics were used to summarise characteristics and outcome data. Continuous variables were considered non-normal if their skewness and kurtosis divided by their standard error exceeded +2 or -2; parametric variables were reported as mean (standard deviation) and non-parametric as median (IQR). Participants were considered “malnourished” if they were rated as SGA B or C, and “well-nourished” if rated SGA A. To determine if there was a significant difference in age, nutrition status, and sex between the samples, the Kruskal-Wallis H Test was applied. If significance was observed, post-hoc testing using Mann-Whitney U was used for pairwise comparisons between cohorts. Here, the Bonferroni correction was applied at 0.05 level and adjusted for three groups; therefore, the cut-off value for significance for Mann-Whitney U tests was at  $p<0.017$ . If significance was observed, post-hoc testing using Mann-Whitney U was used between pairs of cohorts. Post hoc power analysis using G\*Power (version 3.1.9.2) was conducted on independent group means which showed significant difference. Extreme outliers (interquartile

range rule with multiplier of 3) were removed from continuous variables. Binary logistic regression was used to determine the effect of the sample (2012, 2014, or 2019) on nutrition status, with age and sex as a confounding variables.

For the 2014 cohort, differences between nutrition status and outcomes were tested by the Mann-Whitney U Test or Chi-squared test. Associations with health-related outcomes were tested according to the level of severity, with patients considered as “malnourished” (SGA rating B or C), and “severely malnourished” (SGA rating C). Multiple linear regression was used to determine the impact of nutrition status on LOS, and multiple logistic regression was used to determine the impact of nutrition status on secondary outcomes, accounting for participant characteristics which met assumptions for the model. Statistical significance was considered at the  $p<0.05$  level two tailed unless otherwise indicated.

## Results

### Participant characteristics

A total of 286 participants were recruited;  $n=62$  in 2012,  $n=128$  in 2014, and  $n=96$  in 2019 (Table 2). There was no difference in the sex ratio or the prevalence of malnutrition between the sexes in any sample. The 2019 sample (median age 70 years) was found to be significantly younger than the 2012 (median age 80 years) and 2014 (median age 78 years) samples ( $p < 0.017$ ). In the 2014 sample, a circulatory condition was the most common comorbidity experienced (72%), followed by musculoskeletal and respiratory conditions (40% and 38% respectively) (Table 3). Cancer was more prevalent in malnourished (33%) than well-nourished

**Table 3**  
Comorbidity and medication characteristics of 2014 sample

Characteristic	Total sample (n = 124*)	Well-nourished (n= 63)	Malnourished (n= 61)	p-value†
Comorbidities; n (%)				
- Cancer	30 (24)	10 (16)	20 (33)	0.028
- Digestive	38 (31)	22 (35)	16 (26)	0.294
- Musculoskeletal	49 (40)	26 (41)	23 (38)	0.685
- Circulatory	89 (72)	49 (78)	40 (66)	0.131
- Respiratory	47 (38)	22 (35)	25 (41)	0.457
- Nervous	8 (7)	4 (6)	4 (7)	0.962
- Skin	13 (11)	8 (13)	5 (8)	0.413
- Reproductive	11 (9)	5 (8)	6 (10)	0.710
- Kidney	32 (26)	14 (22)	18 (30)	0.354
- Infectious	24 (19)	12 (19)	12 (20)	0.930
- Endocrine	45 (36)	25 (40)	20 (33)	0.425
- Injuries	24 (19)	12 (19)	12 (20)	0.930
- Ear	24 (19)	13 (21)	11 (18)	0.714
- Blood	40 (32)	25 (40)	15 (25)	0.072
- Other	38 (31)	16 (25)	22 (36)	0.198
Cumulative comorbidities‡	4.1 (1.8)	4.2 (1.6)	4.1 (1.9)	0.770
	Total sample (n = 128)	Well-nourished (n= 66)	Malnourished (n= 62)	
Medications class; n (%)				
- Pain	88 (69)	41 (62)	47 (76)	0.095
- Chemotherapy	6 (5)	4 (6)	2 (3)	0.448
- Cardiac	23 (18)	12 (18)	11 (18)	0.948
- Statin	48 (38)	22 (33)	26 (42)	0.315
- Anxiety / depression	40 (31)	19 (29)	21 (34)	0.535
- Hormone/steroid	28 (22)	12 (18)	16 (26)	0.297
- Nausea	17 (13)	6 (9)	11 (18)	0.149
- Nutrition supplement	47 (37)	17 (26)	30 (48)	0.008
- Insomnia	27 (21)	9 (14)	18 (29)	0.033
- Ventolin	41 (32)	17 (26)	24 (39)	0.117
- Insulin	19 (15)	14 (21)	5 (8)	0.037
- Antibiotics	38 (30)	16 (24)	22 (36)	0.164
- Blood pressure	72 (56)	37 (56)	35 (57)	0.964
- Opioid block	4 (3)	1 (2)	3 (5)	0.280
- Constipation	4 (3)	18 (27)	22 (36)	0.317
- PPI§	44 (34)	16 (24)	28 (45)	0.013
- Osteoporosis	14 (11)	8 (12)	6 (10)	0.658
- Incontinence	9 (7)	6 (9)	3 (5)	0.347
- Gout	12 (10)	5 (8)	7 (11)	0.471
- Eyes	15 (12)	9 (14)	6 (10)	0.486
- Epilepsy	18 (14)	10 (15)	8 (13)	0.715
- Blood	51 (40)	26 (39)	25 (40)	0.915
- Parkinson's Disease	6 (5)	3 (5)	3 (5)	0.937
- Allergies	2 (2)	1 (2)	1 (2)	0.964
Cumulative medications	6 (4, 7)	6 (3, 6.25)	6 (5, 7.25)	0.024

\* Comorbidity data in the medical record was unavailable for n=4 participants; † Comparison of well-nourished (SGA rating A) and malnourished (SGA rating B or C) groups; ‡ Number of comorbidities experienced by a single participant; data presented mean (SD); § PPI = proton pump inhibitor; || Number of medication classes taken by a single participant; data presented median (IQR)



**Table 4**  
Health-related outcomes of the 2014 sample according to well-nourished or malnourished

Outcomes	All participants (n=128)	Well-nourished (n=66)	Malnourished (n=62)	P-value*
Length of stay, days	12 (6, 22)	11 (7-19)	13 (6-23)	0.742 †
Hospital readmission	76 (60)	37 (56)	39 (63)	0.431
In hospital mortality	19 (15)	6 (9)	13 (21)	0.059
UTI or respiratory infection	68 (53)	35 (53)	33 (53)	0.982
Falls during admission	38 (30)	20 (30)	18 (29)	0.875
Falls during subsequent admissions	25 (20)	15 (21)	10 (16)	0.347
Pressure wounds	8 (6)	5 (8)	3 (5)	0.523
Fractures on/during admission	22 (17)	12 (18)	10 (16)	0.758

Data expressed as median (IQR) or n (%); \* Comparison of well-nourished (SGA rating A) and malnourished (SGA rating B or C) groups; † Test performed on log10 normalised data

**Table 5**  
Health-related outcomes of the 2014 sample according to severely malnourished or not severely malnourished

Outcomes	No, mild, or moderate malnutrition (n=112)	Severe malnutrition (n=16)	p-value*
Length of stay, days	11.0 (6.0, 20.0)	15.5 (10.5, 31.5)	0.247†
Hospital readmission	69 (62)	7 (44)	0.174
In hospital mortality	13 (12)	6 (38)	0.006
UTI or respiratory infection	61 (55)	7 (44)	0.422
Falls during admission	31 (28)	7 (44)	0.188
Falls during subsequent admissions	25 (22)	0 (0)	0.035
Pressure wounds	8 (7)	0 (0)	0.270
Fractures on/during admission	18 (16)	4 (25)	0.376

Data expressed as median (IQR) or n (%); \* Comparison of no severe malnutrition (SGA rating A or B) and severely malnourished (SGA rating C) groups; † Test performed on log10 normalised data

participants (16%) ( $p=0.028$ ); where prevalence and total number of other comorbidities between groups were similar. Both groups had a median of six classes of medications prescribed during hospitalisation; where the malnourished group had a higher range (IQR 5, 7.25) than the well-nourished group (IQR 3, 6.25) ( $p=0.024$ ). Malnourished participants were also more likely to be prescribed nutritional supplements ( $p=0.008$ ) and proton-pump inhibitors ( $p=0.013$ ), and less likely to be prescribed medication for insomnia ( $p=0.037$ ).

### ***Point-prevalence of malnutrition from 2012 to 2019***

Across the three time-points, malnutrition according to the SGA shows a peak in the 2014 cohort (48%), which was significantly higher than in 2019 (28%), but not 2012 (39%) (Table 2). The prevalence of participants assessed as severely malnourished (SGA rating C) decreased over time from 15% in 2012, 13% in 2014, to 2% in 2019 ( $p=0.005$ ); whereas the prevalence of well-nourished fluctuated from 61% in 2012, 52% in 2014, to 72% in 2019 ( $p=0.005$ ). In a model adjusted for age and sex, regression analyses found that only age was a predictor of

malnutrition, where each year of life increased the odds of malnutrition by 2% (OR: 1.020 [95%CI: 1.003, 1.036]  $p=0.018$ ) but explained only 5% of variation in the model.

'Effect size' and 'chance of impact' was reported for comparisons between year groups for nutrition status to determine the magnitude of the difference between groups and whether or not the outcome was likely to have an actual impact. The pairwise comparison for nutrition status between 2014 and 2019 had a medium effect and 93% chance of impact. For age comparison, 2014 versus 2019 had a medium effect and 89% chance of impact. The 2012 versus 2019 comparison has a small effect size and a 35% chance of impact.

### ***Association of malnutrition with health-related outcomes***

In the 2014 sample, there were five extreme outliers for LOS that were removed. The average LOS was 12 (IQR: 6, 22) days and 60% of participants were readmitted to hospital within 3-months. Malnourished participants (SGA rating B or C) had a higher rate in-hospital mortality but this did not quite reach significance (21% versus 9%;  $p=0.059$ ); groups did not differ on other

outcomes (Table 4). Malnutrition was not a significant predictor of any health-related outcome in adjusted regression models.

Severely malnourished participants (SGA C) had a higher rate in-hospital mortality (38% versus 12%;  $p=0.006$ ), and a lower rate of falls during subsequent admissions (0% versus 22%;  $p=0.035$ ); but groups did not differ on other outcomes (Table 5). Severe malnutrition increased the risk of in-hospital mortality by 457% (OR: 4.57 [95%CI: 1.42, 14.66],  $p=0.011$ ); but in a model adjusted for age, cancer diagnosis, and prescription of nutrition supplements (confounders which met assumptions), this was reduced to 347% with a trend for significance [OR: 3.47 (95%CI: 0.94, 12.78)  $p=0.061$ ]. Severe malnutrition was not a predictor of other health-related outcomes in adjusted or unadjusted models.

## Discussion

This study has reported a consistently high prevalence of malnutrition in three rural hospitals in northern NSW from 2012 to 2019; however, in 2019 the prevalence was 11% and 20% lower than in the previous samples, and the prevalence of severe malnutrition was very low at 2%. The 2014 sample reported the highest prevalence of malnutrition in any Australian hospital (45%); which exceeds the rate reported in three remote Australian hospitals (42%) (19-21). The lowest prevalence reported in 2019 aligns with prevalence rates in Australian metropolitan hospitals. As nutrition status comparison between the 2014 and 2019 cohorts had a calculated medium effect size and high impact value, there is high confidence in the measured prevalence rates. Relevant for the Australian health care system, the SGA tool used to determine prevalence is synonymous with the International Classification for Diseases, 10th revision, Australian Modification (ICD-10-AM) classification of protein-energy malnutrition (22), and therefore directly linked to case-mix funding.

The differing rates in malnutrition prevalence over time is partially explained by the age of participants; however, the impact of age cannot account for the changes in prevalence alone. The lower rate of malnutrition in 2019 compared 2014 may also be due to variations in the sampled wards and hospital sites. The 2019 sample did not include a 34-bed rehabilitation ward; a setting which has previously been reported to have a high prevalence of malnutrition at 53% (23). There are also likely causes of variation in the prevalence of malnutrition over time which were not captured by this study, including demographics such as socio-economic status or ethnicity, inter-rater variability of SGA assessment, or changes in hospital policies and priorities to address malnutrition. Interestingly, the rate of malnutrition did not vary according to sex. A recent meta-analysis of worldwide data found that females had a 45% increased risk of malnutrition (OR: 1.45 [95%CI:

1.27, 1.66]  $p<0.00001$ ) in the community setting, which included post-hospital samples.

It has been well established that malnutrition increases the risk of poor health outcomes in the hospital setting, (21, 24, 25). This study confirms that malnourished participants had higher rates of in-hospital mortality; however, only severe malnutrition was a predictor of this outcome. The clinical importance of this is still relevant despite the rate of severe malnutrition being reduced to only 2% in 2019, as there is a possibility of inter-rater variability in application of the SGA. To confirm if the risk of malnutrition-related in-hospital mortality has been eradicated with the decrease in severe malnutrition rates, health outcomes of the 2019 would need to be examined. A high rate of false positives in the 2014 sample would also explain why many participant characteristics usually associated with malnutrition were not significant predictors in multivariable models. Although severe malnutrition appeared to have a lower rate of falls in subsequent hospital admissions compared to better nourished participants, this is explained by very low rates of hospital readmission reflecting the high rate of in-hospital mortality in this group.

Of clinical significance, less than 50% of malnourished participants in the 2014 sample were provided with nutritional supplementation, and non-supplementation was a predictor of poor health-related outcomes in the adjusted models, including the high rate of hospital readmission within 3-months. Additionally, prescription of proton-pump inhibitors, which inhibit nutrient digestion, was higher in malnourished participants. Overall, the 2014 sample had a high rate of polypharmacy, a risk factor for malnutrition (26). Although comorbidities were highly prevalent, the only disease which was higher in the malnourished participants was cancer. Whilst malnutrition in cancer is known to be highly prevalent (27); previous studies have identified that patients with other hypermetabolic conditions such as hepatic, cardiovascular, and gastrointestinal disease, depression, and dementia also have increased risk of malnutrition (28, 29).

## Limitations

As discussed above, this study is limited by potential poor inter-rater reliability for the SGA assessment between samples, and not including further demographic data to explore variation in the multivariable models. In addition, as the SGA assessments were implemented cross-sectionally, the ratio of pre-existing malnutrition (i.e. admitted to hospital with malnutrition) to hospital acquired malnutrition within the reported prevalence is unclear. Finally, the 2014 sample may have been underpowered to detect differences in health-related outcomes, particularly risk of in-hospital mortality ( $p=0.059$ ).

## Conclusion

Although the prevalence of malnutrition decreased over time, the rate of malnutrition in the sampled rural hospitals was consistently high; and is associated with increased risk of in-hospital mortality. Research should continue to monitor the rate of malnutrition in acute hospitals in rural areas to evaluate the impact of health service policies and procedures to address this problem.

## Key Question Summary

What is known about the topic? Malnutrition is highly prevalent in the acute hospital setting in Australia at 30-40% and up to 71% in older adults (1). The Australian health system faces unique challenges related to high proportions of older adults living in geographically rural and remote areas. The rates and health-related complications of malnutrition in rural and remote Australian hospitals is unexplored.

What does this paper add? This study reported the highest ever recorded prevalence of malnutrition in Australia, at 48% in 2014; which was associated with increased risk of death. However, in 2019 the prevalence has reduced to 28%, and severe malnutrition was almost eradicated (down to 2%).

What are the implications for practitioners? Although the prevalence of malnutrition decreased over time, the rate of malnutrition in the sampled rural hospitals was consistently high; and was associated with increased risk of in-hospital mortality. Less than 50% of malnourished participants were provided with nutritional supplementation, and non-supplementation was a predictor of poor health-related outcomes in adjusted models, including in-hospital mortality.

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**Conflicts of Interest:** MB is the nutrition and dietetics department manager for the sampled sites and oversaw the implementation of all three cross-sectional studies. MB was not involved in the analysis of results. All other authors declare no existing or potential conflict of interest.

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